KU

Definition of the Simple Regression Model

Deriving the Ordinary Leas Squares Estimates

Properties of OLS on any Sample of Data

Units of Measurement and Function Form

Using the Natural Logarithm in Simp Regression

Expected Value of OLS

### The Simple Regression Model

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These slides were based on Introductory Econometrics by Jeffrey M. Wooldridge (2015)

# KU Topics

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### 1 Definition of the Simple Regression Model

- ② Deriving the Ordinary Least Squares Estimates
- <sup>(3)</sup> Properties of OLS on any Sample of Data

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### KU Definition of the Simple Regression Model

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- What type of analysis will we do? Cross-sectional analysis
- **First step:** Clearly define what is your population (in what you are interested to study).
- Second Step: There are two variables, x and y, and we would like to "study how y varies with changes in x."
- **Third Step:** We assume we can collect a random sample from the population of interest.

Now we will learn to write our first econometric model, derive an estimator (what's an estimator again?) and use this estimator in our sample.

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### We must confront three issues:

- **1** How do we allow factors other than x to affect y? There is never an exact relationship between two variables.
- **2** What is the functional relationship between y and x?
- **3** How can we be sure we are capturing a *ceteris paribus* relationship between y and x?

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Expected Value of OLS Consider the following equation relating y to x:

 $y = \beta_0 + \beta_1 x + u,$ 

which is assumed to hold in the population of interest.

• This equation defines the **simple linear regression model** (or *two-variable regression model*, *or bivariate linear regression model*).

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Expected Value of OLS • y and x are not treated symmetrically. We want to explain y in terms of x.

 $x ext{ explains } y$  $x \longrightarrow y$ 

• Example:

size of the city x, explains number of crimes (y) (not the other way around).

### KU Terminology for Simple Regression

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У	x
Dependent Variable	Independent Variable
Explained Variable	Explanatory Variable
Response Variable	Control Variable
Predicted Variable	Predictor Variable
Regressand	Regressor

### $K\!U$ The error term

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Expected Value of OL  $y = \beta_0 + \beta_1 x + u$ 

This equation explicitly allows for other factors, contained in u, to affect y.

This equation also addresses the functional form issue (in a simple way).

Namely, y is assumed to be *linearly* related to x.

We call  $\beta_0$  the **intercept parameter** and  $\beta_1$  the **slope parameter**. These describe a population, and our ultimate goal is to estimate them.

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Expected Value of OLS • The equation also addresses the *ceteris paribus* issue. In

$$y = \beta_0 + \beta_1 x + u,$$

all other factors that affect y are in u. We want to know how y changes when x changes, *holding* u *fixed*.

• Let  $\Delta$  denote "change."Then holding u fixed means  $\Delta u=0.$  So

$$\begin{array}{rcl} \Delta y &=& \beta_1 \Delta x + \Delta u \\ &=& \beta_1 \Delta x & \mbox{ when } \Delta u = 0. \end{array}$$

• This equation effectively defines  $\beta_1$  as a slope, with the only difference being the restriction  $\Delta u = 0$ .

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#### Example: Yield and Fertilizer

• A model to explain crop yield to fertilizer use is

 $yield = \beta_0 + \beta_1 fertilizer + u,$ 

where u contains land quality, rainfall on a plot of land, and so on. The slope parameter,  $\beta_1$ , is of primary interest: it tells us how *yield* changes when the amount of fertilizer changes, holding all else fixed.

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#### **Example:** Wage and Education

$$wage = \beta_0 + \beta_1 educ + u$$

where u contains somewhat nebulous factors ("ability") but also past workforce experience and tenure on the current job.

 $\Delta wage = \beta_1 \Delta educ \quad \text{ when } \Delta u = 0$ 

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Expected Value of OL We said we must confront three issues:

1. How do we allow factors other than x to affect y?

Answer: u

2. What is the functional relationship between y and x? **Answer:** Linear (x has a linear effect on y)

3. How can we be sure we a capturing a ceteris paribus relationship between  $\boldsymbol{y}$  and  $\boldsymbol{x}?$ 

**Answer:** Related with  $\Delta u = 0$ 

• We have argued that the simple regression model

$$y = \beta_0 + \beta_1 x + u$$

addresses each of them.

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Expected Value of OLS To estimate  $\beta_1$  and  $\beta_0$  from a random sample we also need to restrict how u and x are related to each other.

- Recall that x and u are properly viewed as having distributions in the population.
- What we must do is restrict the way in when u and x relate to each other in the **population**.
- First, we make a simplifying assumption that is without loss of generality: the average, or expected, value of u is zero in the population:

$$E(u) = 0$$

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Expected Value of OL

- Normalizing u should cause no impact in the most important parameter:  $\beta_1$
- The presence of  $\beta_0$  in

$$y = \beta_0 + \beta_1 x + u$$

allows us to assume E(u) = 0.

• If the average of u is different from zero, we just adjust the intercept, leaving the slope the same. If  $\alpha_0 = E(u)$  then we can write

$$y = (\beta_0 + \alpha_0) + \beta_1 x + (u - \alpha_0),$$

where the new error,  $u - \alpha_0$ , has a zero mean.

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Expected Value of OLS We need to restrict the dependence between u and x

### • Option 1: Uncorrelated

We could assume u and x **uncorrelated** in the population:

Corr(x, u) = 0

It implies only that u and x are not linearly related. Not good enough.

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### • **Option 2:** Mean independence

The mean of the error (i.e., the mean of the unobservables) is the same across all slices of the population determined by values of x.

We represent it by:

$$E(u|x) = E(u)$$
, all values  $x$ ,

And we say that u is **mean independent** of x

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Expected Value of OLS • Suppose u is "ability" and x is years of education. We need, for example,

$$E(ability|x = 8) = E(ability|x = 12) = E(ability|x = 16)$$

so that the average ability is the same in the different portions of the population with an  $8^{th}$  grade education, a  $12^{th}$  grade education, and a four-year college education.

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Expected Value of OLS  $\bullet$  Combining E(u|x)=E(u) (the substantive assumption) with E(u)=0 (a normalization) gives

$$E(u|x) = 0$$
, all values  $x$ 

• Called the zero conditional mean assumption.

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- First, recall the properties of conditional expectation. (see slides with a review of *Probability*)
- Now, take the conditional expectation of our *Simple Linear Regression Function*. Then, we get:

$$E(y|x) = \beta_0 + \beta_1 x + E(u|x)$$
  
=  $\beta_0 + \beta_1 x$ 

which shows the **population regression function** is a linear function of x.

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### Figure: Example: The goal is to explain weekly consumption expenditure in terms of weekly income

$Y_{\downarrow}$ $X \rightarrow$	80	100	120	140	160	180	200	220	240	26
Weekly family	55	65	79	80	102	110	120	135	137	150
consumption	60	70	84	93	107	115	136	137	145	15
expenditure Y, \$	65	74	90	95	110	120	140	140	155	175
	70	80	94	103	116	130	144	152	165	178
	75	85	98	108	118	135	145	157	175	180
	-	88	-	113	125	140	-	160	189	185
	-	-	-	115	-	-	-	162	-	19
Total	325	462	445	707	678	750	685	1043	966	121
Conditional means of $Y$ , $E(Y X)$	65	77	89	101	113	125	137	149	161	17

Source: Gujarati, Damodar (2002). Basic Econometrics.

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#### Figure: Conditional Probabilities of the data

$X \rightarrow$										
$p(Y \mid X_i)$	80	100	120	140	160	180	200	220	240	260
Conditional probabilities $p(Y   X_i)$	$\frac{1}{5}$	$\frac{1}{6}$	$\frac{1}{5}$	$\frac{1}{7}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{5}$	$\frac{1}{7}$	$\frac{1}{6}$	$\frac{1}{7}$
	1 5	$\frac{1}{6}$	1 5	$\frac{1}{7}$	$\frac{1}{6}$	$\frac{1}{6}$	1 5	$\frac{1}{7}$	$\frac{1}{6}$	1 7
	1 5	$\frac{1}{6}$	1 5	$\frac{1}{7}$	1 6	1 6	1 5	$\frac{1}{7}$	1 6	$\frac{1}{7}$
	$\frac{1}{5}$	$\frac{1}{6}$	<u>1</u> 5	$\frac{1}{7}$	$\frac{1}{6}$	$\frac{1}{6}$	<u>1</u> 5	$\frac{1}{7}$	$\frac{1}{6}$	$\frac{1}{7}$
	$\frac{1}{5}$	$\frac{1}{6}$	$\frac{1}{5}$	$\frac{1}{7}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{5}$	$\frac{1}{7}$	$\frac{1}{6}$	$\frac{1}{7}$
	-	$\frac{1}{6}$	-	$\frac{1}{7}$	$\frac{1}{6}$	$\frac{1}{6}$	-	$\frac{1}{7}$	$\frac{1}{6}$	$\frac{1}{7}$
	-	-	-	$\frac{1}{7}$	-	-	-	$\frac{1}{7}$	-	$\frac{1}{7}$
Conditional means of <i>Y</i>	65	77	89	101	113	125	137	149	161	173

Source: Gujarati, Damodar (2002). Basic Econometrics.

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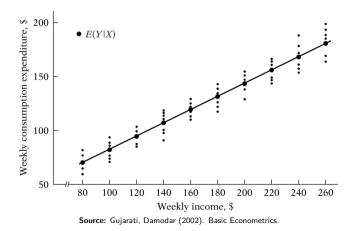
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#### Figure: Conditional distribution of expenditure for various levels of income



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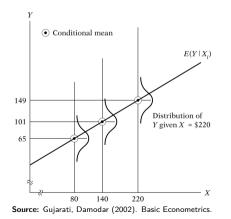
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#### Figure: The Population Regression Function (PRF)



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- The straight line in the previous graph is the PRF,  $E(y|x) = \beta_0 + \beta_1 x$ . The conditional distribution of y at three different values of x are superimposed.
- For a given value of x, we see a range of y values: remember,  $y = \beta_0 + \beta_1 x + u$ , and u has a distribution in the population.
- In practice, we never know the **population intercept and slope**.

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Expected Value of OLS • Assuming we know the PRF, consider this example:

#### Example

• Suppose for the population of students attending a university, we know the PRF:

E(colGPA|hsGPA) = 1.5 + 0.5 hsGPA,

- For this example, what is y? what is x? What is the slope? What's the intercept?
- If hsGPA = 3.6 what's the expected college GPA? 1.5 + 0.5(3.6) = 3.3

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# $\displaystyle{K\!U}$ Deriving the OLS

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Expected Value of OLS • Given data on x and y, how can we estimate the population parameters,  $\beta_0$  and  $\beta_1?$ 

• Let  $\{(x_i, y_i) : i = 1, 2, ..., n\}$  be a **random sample** of size n (the number of observations) from the population. Think of this as a random sample.

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### Derivation: (On white board)

### Estimator for $\beta_0$

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$

### Estimator for $\beta_1$

$$\begin{split} \hat{\beta}_1 &= \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\text{Sample Covariance}(x, y)}{\text{Sample Variance}(x)} \\ &= \frac{S_{x,y}}{S_x^2} \\ &= \hat{\rho}_{x,y} \frac{\hat{\sigma}_y}{\hat{\sigma}_x} \end{split}$$

# $\displaystyle{K\!U}$ Deriving the OLS

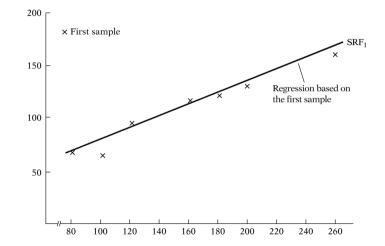
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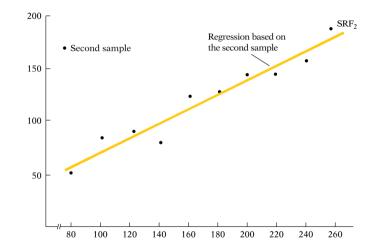
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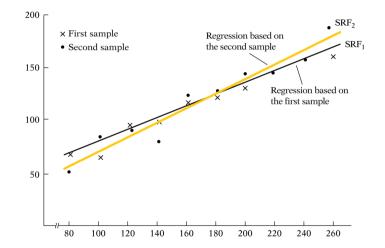
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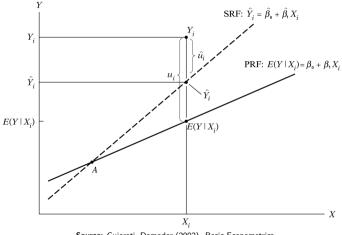
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#### Example: Effects of Education on Hourly Wage

• Data: random sample from the US workforce population in 1976. *wage*: dollars per hour, *educ*: highest grade completed (years of education).

• The estimated equation is

$$\widehat{wage} = -0.90 + 0.54 \ educ$$
$$n = 526$$

• Each additional year of schooling is estimated to be worth \$0.54.

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### The function

$$\widehat{wage} = -0.90 + 0.54 \ educ$$

### is the OLS (or sample) regression line.

### KU Interpreting the OLS Estimates - R Output

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> stargazer(regress	ion_wage1,	<pre>type='text',</pre>	align=TRUE,	digits=2)						
Dependent variable:										
		wage								
educ		0.54*** (0.05)								
Constant		-0.90 (0.68)								
Observations R2		526 0.16								
Adjusted R2 Residual Std. Error F Statistic		0.16 8 (df = 524) ** (df = 1; 52	24)							
Note:		*p<0.05; ***p								

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Expected Value of OLS • When we write the simple linear regression model,

$$wage = \beta_0 + \beta_1 educ + u,$$

- it applies to the population, so we do not know  $\beta_0$  and  $\beta_1$ .
- $\hat{\beta}_0 = -0.90$  and  $\hat{\beta}_1 = 0.54$  are our *estimates* from this particular sample.
- These estimates may or may not be close to the population values. If we obtain another sample, the estimates would almost certainly change.

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Expected Value of OLS • If educ = 0, the predicted wage is:

$$\widehat{wage} = -0.90 + 0.54(0) = -0.90$$

The predicted value does not fit in reality.

Mainly because when we extrapolate outside the majority range of our data can produce strange predictions.

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Expected Value of OLS • When educ = 8, the predicted wage is:

$$\widehat{wage} = -0.90 + 0.54(8) = 3.42$$

which we can think of as our estimate of the average wage in the population when educ = 8.

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### Sample Regression Line (SRF)

$$\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i \qquad i = 1, \dots, n$$

Also known as:

- OLS Regression Line
- Sample Regression Function
- OLS Regression Function
- Estimated Equation

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### **Population Regression Function (PRF)**

Since the simple linear regression model (or just econometric model) is:

$$y_i = \beta_0 + \beta_1 x_i + u$$

Then, the PRF is:

$$\Rightarrow E(y_i|x) = \beta_0 + \beta_1 x_i \qquad i = 1, 2, \dots, n$$

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### Residuals

$$\hat{u}_i = y_i - \hat{y}_i \qquad i = 1, 2, \dots, n$$

### **Error Term**

$$u_i = y_i - E(y|x)$$
  
=  $y_i - \beta_0 - \beta_1 x_i$   $i = 1, 2, ..., n$ 

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### **5** Expected Value of OLS

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Expected Value of OLS • Recall that the OLS residuals are

$$\hat{u}_i = y_i - \hat{y}_i = y_i - \hat{eta}_0 - \hat{eta}_1 x_i$$
 ,  $i = 1, 2, ..., n$ 

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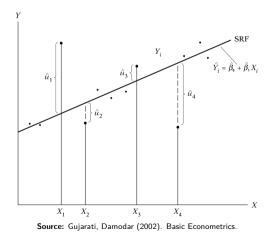
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- Some residuals are positive, others are negative.
- If  $\hat{u}_i$  is positive  $\Rightarrow$  the line underpredicts  $y_i$
- If  $\hat{u}_i$  is negative  $\Rightarrow$  the line overpredicts  $y_i$

## KU Algebraic Properties of OLS Statistics

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### (1) The sum of the OLS residuals is 0

$$\sum_{i=1}^{n} \hat{u}_i = 0$$

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Expected Value of OLS  $(\mathbf{2})$  The sample covariance between the explanatory variables and the residuals is always zero

$$\sum_{i=1}^{n} x_i \hat{u}_i = 0$$

- Therefore the sample correlation between the x and  $\hat{u}_i$  is also equal to zero.
- Because the  $\hat{y}_i$  are linear functions of the  $x_i$ , the fitted values and residuals are uncorrelated, too:

$$\sum_{i=1}^{n} \hat{y}_i \hat{u}_i = 0$$

## KU Algebraic Properties of OLS Statistics

Definition of the Simple Regression Model

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Expected Value of OLS

### (3) The point $(\bar{x}, \bar{y})$ is always on the OLS regression line.

$$\bar{y} = \hat{\beta}_0 + \hat{\beta}_1 \bar{x}$$

• That is, if we plug in the average for x, we predict the sample average for y.



Definition of the Simple Regression Model

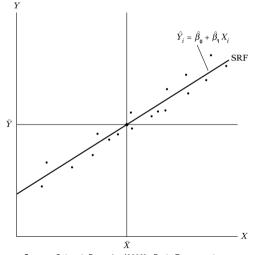
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Source: Gujarati, Damodar (2002). Basic Econometrics.

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### Goodness-of-Fit

• For each observation, write

$$y_i = \hat{y}_i + \hat{u}_i$$

• Define:

Total Sum of Squares  $= SST = \sum_{i=1}^{n} (y_i - \bar{y})^2$ Explained Sum of Squares  $= SSE = \sum_{i=1}^{n} (\hat{y}_i - \bar{y})^2$ Residual Sum of Squares  $= SSR = \sum_{i=1}^{n} \hat{u}_i^2$ 

KU Goodness-of-Fit

Definition of the Simple Regression Model

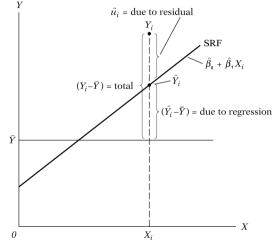
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Source: Gujarati, Damodar (2002). Basic Econometrics.

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### (Other names)

- SSR is also know as *Sum of Squared Residuals* or *Model Sum of Residuals*
- SST = TSS
- SSE = ESS
- SSR = RSS

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$$SST = \sum_{i=1}^{n} (y_i - \bar{y})^2$$
$$= \sum_{i=1}^{n} [(y_i - \hat{y}_i) + (\hat{y}_i - \bar{y})]^2$$
$$= \sum_{i=1}^{n} [\hat{u}_i - (\hat{y}_i - \bar{y})]^2$$

Using the fact that the fitted values and residuals are uncorrelated:

$$SST = SSE + SSR$$

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### The R-Squared

**Goal:** We want to evaluate how well the independet variable x explains the dependent variable y.

- We want to obtain the fraction of the sample variation in y that is explained by x.
- We will summarize it in one number:  $R^2$  (or coefficient of determination.)
- Assuming SST > 0,

$$R^2 = \frac{SSE}{SST} = 1 - \frac{SSR}{SST}$$

# KU Goodness-of-Fit

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Expected Value of OLS • Since SSE cannot be greater than the SST, then:

$$0 \le R^2 \le 1$$

- $R^2 = 0 \Rightarrow$  No linear relationship (between  $y_i$  and  $x_i$ ).
- $R^2 = 1 \Rightarrow$  **Perfect linear relationship** (between  $y_i$  and  $x_i$ ).
- As  $R^2$  increases  $\Rightarrow y_i$  gets closer and closer to the OLS regression line.

We should not focus only on  $\mathbb{R}^2$  to analyze our regression.

# KU Goodness-of-Fit

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### **Example** (Wage)

$$\widehat{wage} = -0.90 + 0.54 \ educ$$
  
 $n = 526, \quad R^2 = .16$ 

 $\bullet$  Therefore, years of education explains only about 16% of the variation in hourly wage.

# KU Goodness-of-Fit - R output

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```
> summary(regression_wage1)
```

```
Call:
lm(formula = wage ~ educ, data = wage1)
```

Residuals:

Min	1Q	Median	ЗQ	Max
-5.3396	-2.1501	-0.9674	1.1921	16.6085

```
Coefficients:
```

Estimate Std. Error t value Pr(>|t|) (Intercept) -0.90485 0.68497 -1.321 0.187 educ 0.54136 0.05325 10.167 <2e-16 \*\*\* ---Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.378 on 524 degrees of freedom Multiple R-squared: 0.1648, Adjusted R-squared: 0.1632 F-statistic: 103.4 on 1 and 524 DF, p-value: < 2.2e-16

### **KU** Goodness-of-Fit - R output (using *stargazer*)

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> stargazer(regression_wage1, type='text', align=TRUE, digits=2				
	-	dent variable		
		wage		
educ		0.54*** (0.05)		
Constant		-0.90 (0.68)		
Observations R2 Adjusted R2 Residual Std. Error F Statistic		526 0.16 0.16 8 (df = 524) ** (df = 1; 52	-	
Note:	*p<0.1; *	*p<0.05; ***p		

# KU Suggested Exercise

Definition of the Simple Regression Model

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### **Suggested Exercise**

Below you have a random sample with 10 data points. Your observations are  $(x_i, y_i)$ . Find the  $\hat{\beta}_0$ ,  $\hat{\beta}_1$  and the  $R^2$ .

Obs. #	$x_i$	$y_i$	$x_i$	$(y_i - ar y)$	$(x_i - \bar{x})$	$(y_i - ar y)^2$	$(x_i - \bar{x})^2$	$(x_i-ar x)(y_i-ar y)$	$\hat{y}_i$	$(y_i - \hat{y}_i)$	$(\hat{y}_i - \bar{y})^2$	$(y_i - \hat{y}_i)^2$
1	$x_1$	70	80	-41	-90	1681	8100	3690	65.18	4.82	2099.31	23.21
2	$x_2$	65	100	-46	-70	2116	4900	3220	75.36	-10.36	1269.95	107.40
3	$x_3$	90	120	-21	-50	441	2500	1050	85.55	4.45	647.93	19.84
4	$x_4$	95	140	-16	-30	256	900	480	95.73	-0.73	233.26	0.53
5	$x_5$	110	160	-1	-10	1	100	10	105.91	4.09	25.92	16.74
6	$x_6$	115	180	4	10	16	100	40	116.09	-1.09	25.92	1.19
7	$x_7$	120	200	9	30	81	900	270	126.27	-6.27	233.26	39.35
8	$x_8$	140	220	29	50	841	2500	1450	136.45	3.55	647.93	12.57
9	$x_9$	155	240	44	70	1936	4900	3080	146.64	8.36	1269.95	69.95
10	$x_{10}$	150	260	39	90	1521	8100	3510	156.82	-6.82	2099.31	46.49
	$\mathbf{Sum}$	1,110	1,700	0.00	0.00	8,890	33,000	16,800	1,110	0.00	$^{8,553}$	337

# KU Topics

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### Units of Measurement and Functional Form

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- **2** Deriving the Ordinary Least Squares Estimates
- <sup>3</sup> Properties of OLS on any Sample of Data

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### **5** Expected Value of OLS

# ${f K}{f U}$ The effects of Changing Units of Measurement on OLS Statistics

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### Example

salary: Annual CEO's salary in thousands of dollars roe: Average return on equity (measured in percentage)

$$\widehat{salary} = 963.19 + 18.50 \ root$$
  
 $n = 209, \ R^2 = .01$ 

• A one unit increase in the independent variable (i.e. roe increases one percent)  $\Rightarrow$  increases the predicted salary by 18.501, or **\$18,501**.

# ${ m K\!U}$ The effects of Changing Units of Measurement on OLS Statistics

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Expected Value of OLS If we measure *roe* as a decimal (rather than a percent), what will happen to the intercept, slope, and R<sup>2</sup>?
 We want:

roedec = roe/100

• What if we measure salary in dollars (rather than thousands of dollars)? what will happen to the intercept, slope, and  $R^2$ ? We want:

 $salarydol = 1,000 \cdot salary$ 

# ${f K}{f U}$ The effects of Changing Units of Measurement on OLS Statistics

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### Changing Units of Measurement

• If the dependent variable y is multiplied by a constant  $c\Rightarrow c\cdot\hat{\beta}_0$  and  $c\cdot\hat{\beta}_1$ 

• If the independent variable x is multiplied by a constant  $c \Rightarrow \frac{1}{c} \cdot \hat{\beta}_1$ 

In general, changing the units of measurement of only the independent variable does not affect the intercept

# $K\!U$ The effects of Changing Units of Measurement on OLS Statistics

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### **Example:** CEO's salary - Original Regression

$$\widehat{salary} = 963.19 + 18.50 \ roe$$
  
 $n = 209, \ R^2 = .01$ 

### **Example:** CEO's salary - *roe* as a decimal

The new regression is:

$$\widehat{salary} = 963.191 + 1,850.1 \ roedec$$
  
 $n = 209, \ R^2 = .01$ 

### KU The effects of Changing Units of Measurement on OLS Statistics

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Expected Value of OLS > roedec<-ceosal1\$roe\*(1/100)</pre>

> regression\_ceosal1c <- lm(salary ~ roedec, data = ceosal1)</pre>

> stargazer(regression\_ceosal1c, type='text', align=TRUE, digits=2)

Dependent variable:

\_\_\_\_\_

salary

-----

roedec	1,850.12*
	(1,112.33)

Constant	963.19***
	(213.24)

# $K\!U$ The effects of Changing Units of Measurement on OLS Statistics

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Expected Value of OLS

### Example: CEO's salary - Original Regression

$$\widehat{salary} = 963.19 + 18.50 \ roe$$
  
 $n = 209, \ R^2 = .01$ 

### **Example:** CEO's salary - *salary* in dollars

The new regression is

$$\widehat{salarydol} = 963, 191 + 18, 501 \ roe$$
  
 $n = 209, \ R^2 = .01$ 

### ${f KU}$ The effects of Changing Units of Measurement on OLS Statistics

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Expected Value of OLS

<pre>&gt; stargazer(regression_ceosal1b, type='text', align=TRUE, dig </pre>	<pre>&gt; salarydol&lt;-ceosal1 &gt; regression_ceosal1</pre>	b <- lm(salarydol ~ roe, data = ceosal1)
Dependent variable:           salarydol           roe         18,501.19*           (11,123.25)           Constant         963,191.30***           (213,240.30)           Observations         209           R2         0.01           Adjusted R2         0.01           Residual Std. Error         1,366,555.00 (df = 207)		
salarydol           roe         18,501.19*           (11,123.25)           Constant         963,191.30***           (213,240.30)           Observations         209           R2         0.01           Adjusted R2         0.01           Residual Std. Error         1,366,555.00 (df = 207)		
roe 18,501.19* (11,123.25) Constant 963,191.30*** (213,240.30) 		Dependent variable:
(11,123.25) Constant 963,191.30*** (213,240.30)  Observations 209 R2 0.01 Adjusted R2 0.01 Residual Std. Error 1,366,555.00 (df = 207)		salarydol
Constant 963,191.30*** (213,240.30) Dbservations 209 R2 0.01 Adjusted R2 0.01 Residual Std. Error 1,366,555.00 (df = 207)	roe	18,501.19*
(213,240.30) Observations 209 R2 0.01 Adjusted R2 0.01 Residual Std. Error 1,366,555.00 (df = 207)		(11,123.25)
Observations         209           R2         0.01           Adjusted R2         0.01           Residual Std. Error         1,366,555.00 (df = 207)	Constant	963,191.30***
R2         0.01           Adjusted R2         0.01           Residual Std. Error         1,366,555.00 (df = 207)		(213,240.30)
R2 0.01 Adjusted R2 0.01 Residual Std. Error 1,366,555.00 (df = 207)	Observations	209
Residual Std. Error 1,366,555.00 (df = 207)	R2	
	Adjusted R2	0.01
F Statistic 2.77* (df = 1: 207)	Residual Std. Error	1,366,555.00 (df = 207)
	F Statistic	2.77* (df = 1; 207)

Definition of the Simple Regression Model

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Expected Value of OLS • Recall the **wage** example:

### Example (Wage)

$$\widehat{vage} = -0.90 + 0.54 \ educ$$
  
 $n = 526, \quad R^2 = .16$ 

- $\bullet$  Now, think about the econometric model and how this OLS Regression Function is interpreted.
- What the OLS Regression Line says may not fit how economically we see the problem.

Possible issue: the dollar value of another year of schooling is constant.

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- Expected Value of OLS

- So the  $16^{th}$  year of education is worth the same as the second.
  - We expect additional years of schooling to be worth more, in dollar terms, than previous years.
- How can we incorporate an increasing effect? One way is to postulate a constant *percentage* effect.
- •We can approximate percentage changes using the natural log.

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### **Constant Percent Model**

• Let the dependent variable be  $\log(wage)$  and write a (new) simple linear regression model:

$$\log(wage) = \beta_0 + \beta_1 educ + u$$

• Let's define log(wage) (write it as lwage) and run a new regression.

### $\ensuremath{{\ensuremath{\mathsf{KU}}}}$ Using the Natural Logarithm in Simple Regression

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	Dependent variable:
	lwage
educ	0.08***
	(0.01)
Constant	0.58***
	(0.10)
Observations	526
R2	0.19
Adjusted R2	0.18
Residual Std. Error	0.48 (df = 524)
F Statistic	119.58*** (df = 1; 524)
Note:	*p<0.1; **p<0.05; ***p<0.01

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Expected Value of OLS

$$\widehat{wage} = 0.58 + .08 \ educ$$
  
 $n = 526, R^2 = .19$ 

• The estimated return to each year of education is about 8%.

### • Attention:

This R-squared is not directly comparable to the R-squared when wage is the dependent variable. The total variation (SSTs) in  $wage_i$  and  $lwage_i$  that we must explain are completely different.

### KU Using the Natural Logarithm in Simple Regression

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#### **Constant Elasticity Model**

• We can use the log on both sides of the equation to get **constant elasticity models**. For example, if

 $\log(salary) = \beta_0 + \beta_1 \log(sales) + u$ 

then

$$\beta_1 \approx \frac{\% \Delta salary}{\% \Delta sales}$$

- The elasticity is free of units of *salary* and *sales*.
- A constant elasticity model for salary and sales makes more sense than a constant dollar effect.

### $K\!U$ Using the Natural Logarithm in Simple Regression

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Model	Dependent Variable	Independent Variable	Interpretation of $eta_1$
Level-Level	y	x	$\Delta y = \beta_1 \Delta x$
Level-Log	y	$\log(x)$	$\Delta y = (\beta_1/100)\%\Delta x$
Log-Level	$\log(y)$	x	$\%\Delta y = (100\beta_1)\Delta x$
Log-Log	$\log(y)$	$\log(x)$	$\%\Delta y = \beta_1\%\Delta x$

### KU Using the Natural Logarithm in Simple Regression

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Expected Value of OLS • Recall the CEO salary example, but now the independent variable is sales.

$$salary = \beta_o + \beta_1 sales + u$$

• Applying log on both variables (dependent and independent) we get:

#### **Example** (CEO salary)

$$\widehat{log(salary)} = 4.82 + 0.26 \ log(salars)$$
$$n = 209, \qquad R^2 = .21$$

#### $\displaystyle K\!U$ Using the Natural Logarithm in Simple Regression

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Expected Value of OLS

	Dependent variable:	
-	log(salary)	
log(sales)	0.26*** (0.03)	
Constant	4.82*** (0.29)	
Observations R2 Adjusted R2 Residual Std. Error F Statistic	209 0.21 0.21 0.50 (df = 207) 55.30*** (df = 1; 207)	
Note: *	p<0.1; **p<0.05; ***p<0.01	

### KU Using the Natural Logarithm in Simple Regression

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- The estimated elasticity of CEO salary with respect to firms sales is about .26.
  - A 10 percent increase in sales is associated with a

.26(10) = 2.6

percent increase in salary.

# KU Topics

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#### **5** Expected Value of OLS

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Expected Value of OLS

#### Goal: We want to study statistical properties of the OLS estimator

• In order to that, we will need to impose 4 assumptions.

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Expected Value of OLS

#### Assumption SLR.1 (Linear in Parameters)

The population model can be written as

$$y = \beta_0 + \beta_1 x + u$$

where  $\beta_0$  and  $\beta_1$  are the (unknown) population parameters.

- What linear in parameters mean?
- Example of non linear in parameters on white board

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#### Assumption SLR.2 (Random Sampling)

We have a random sample of size n,  $\{(x_i, y_i) : i = 1, ..., n\}$ , following the population model.

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#### Assumption SLR.3 (Sample Variation in the Explanatory Variable)

The sample outcomes on  $x_i$  are not all the same value.

- This is the same as saying the sample variance of  $\{x_i : i = 1, ..., n\}$  is **not zero**.
- $\bullet$  If in the population x does not change then we are not asking an interesting question.

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#### Assumption SLR.4 (Zero Conditional Mean)

In the population, the error term has zero mean given any value of the explanatory variable:

$$E(u|x) = 0$$
 for all  $x$ .

#### • Key assumption.

• We can compute the OLS estimates whether or not this assumption holds.

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• If.

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Expected Value of OLS **Goal:** We want to know if  $\hat{\beta}_1$  is unbiased for  $\beta_1$ , and  $\hat{\beta}_0$  is unbiased for  $\beta_0$ 

 $E(\hat{\beta}_1) = \beta_1$ 

 $E(\hat{\beta}_0) = \beta_0$ 

Then, the **OLS estimator** is unbiased.

• **Demonstration:** On the white board.

## KU Unbiasedness of OLS

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Expected Value of OLS

#### Theorem: Unbiasedness of OLS

Under Assumptions SLR.1 through SLR.4

$$E(\hat{eta}_0)=eta_0$$
 and  $E(\hat{eta}_1)=eta_1$  ,

for any values of  $\beta_0$  and  $\beta_1$ , i.e.,  $\hat{\beta}_0$  is unbiased for  $\beta_0$ , and  $\hat{\beta}_1$  is unbiased for  $\beta_1$ 

# $\displaystyle{K\!U}$ Unbiasedness of OLS

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Expected Value of OLS • Therefore, the four assumptions for the OLS estimator to be unbiased are:

**SLR.1:** (Linear in Parameters)  $y = \beta_0 + \beta_1 x + u$ **SLR.2:** (Random Sampling) **SLR.3:** (Sample Variation in  $x_i$ ) **SLR.4:** (Zero Conditional Mean) E(u|x) = 0

• If any of these assumptions fails, the OLS estimator will (generally) be biased.

• To be discussed in the next chapter: What are the omitted factors? Are they likely to be correlated with x? If so, SLR.4 fails and OLS will be biased.